



# Introduction to Data Science

**GIRI NARASIMHAN, SCIS, FIU**

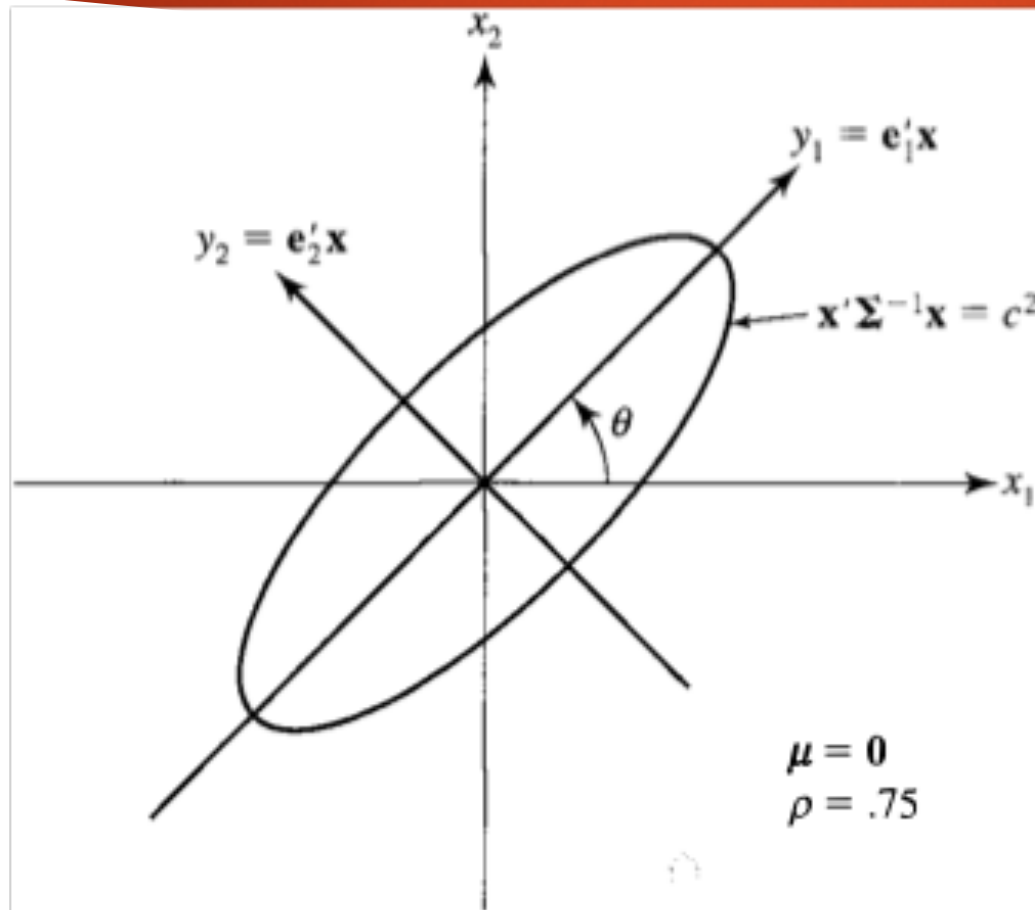
# PCA and Matrices

**FROM JOHNSON & WICHERN, *APPLIED  
MULTIVARIATE STATISTICAL ANALYSIS*, 6TH ED**

# PCA: Principal Component Analysis

- ▶ Tool for Dimensionality Reduction
  - ▣ Reduces impact of curse of dimensionality
- ▶ Tool for finding Subspace in which data lies
- ▶ Summarization of data to find important variables
- ▶ Compares relative importance of variables
- ▶ Explains the most amount of variation in data

# Principal Components



**Figure 8.1** The constant density ellipse  $\mathbf{x}' \boldsymbol{\Sigma}^{-1} \mathbf{x} = c^2$  and the principal components  $y_1, y_2$  for a bivariate normal random vector  $\mathbf{X}$  having mean  $\mathbf{0}$ .

# Principal Components

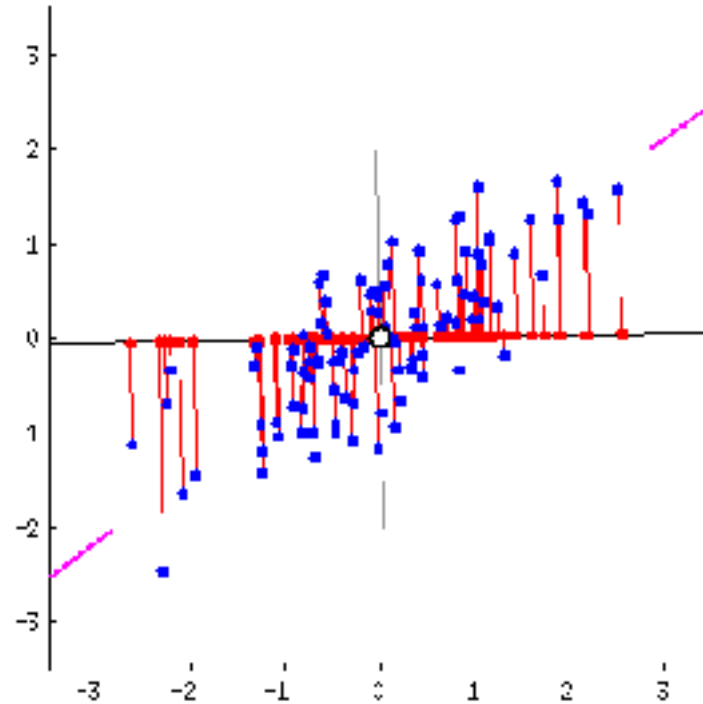
First *sample* principal component = linear combination  $\mathbf{a}'_1 \mathbf{x}_j$  that maximizes the sample variance of  $\mathbf{a}'_1 \mathbf{x}_j$  subject to  $\mathbf{a}'_1 \mathbf{a}_1 = 1$

Second *sample* principal component = linear combination  $\mathbf{a}'_2 \mathbf{x}_j$  that maximizes the sample variance of  $\mathbf{a}'_2 \mathbf{x}_j$  subject to  $\mathbf{a}'_2 \mathbf{a}_2 = 1$  and zero sample covariance for the pairs  $(\mathbf{a}'_1 \mathbf{x}_j, \mathbf{a}'_2 \mathbf{x}_j)$

At the *i*th step, we have

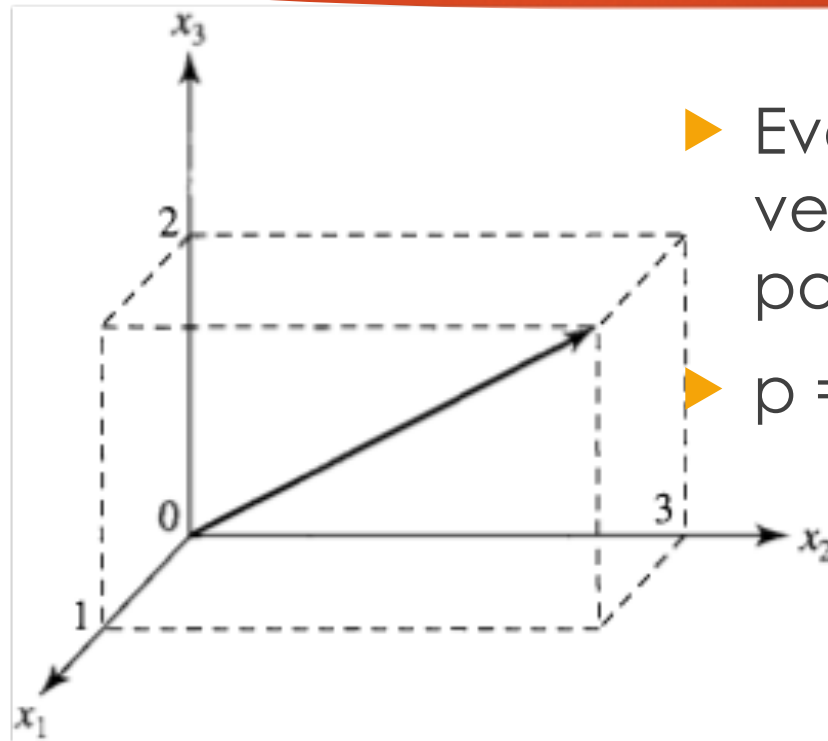
*i*th *sample* principal component = linear combination  $\mathbf{a}'_i \mathbf{x}_j$  that maximizes the sample variance of  $\mathbf{a}'_i \mathbf{x}_j$  subject to  $\mathbf{a}'_i \mathbf{a}_i = 1$  and zero sample covariance for all pairs  $(\mathbf{a}'_i \mathbf{x}_j, \mathbf{a}'_k \mathbf{x}_j)$ ,  $k < i$

# PCA Animation



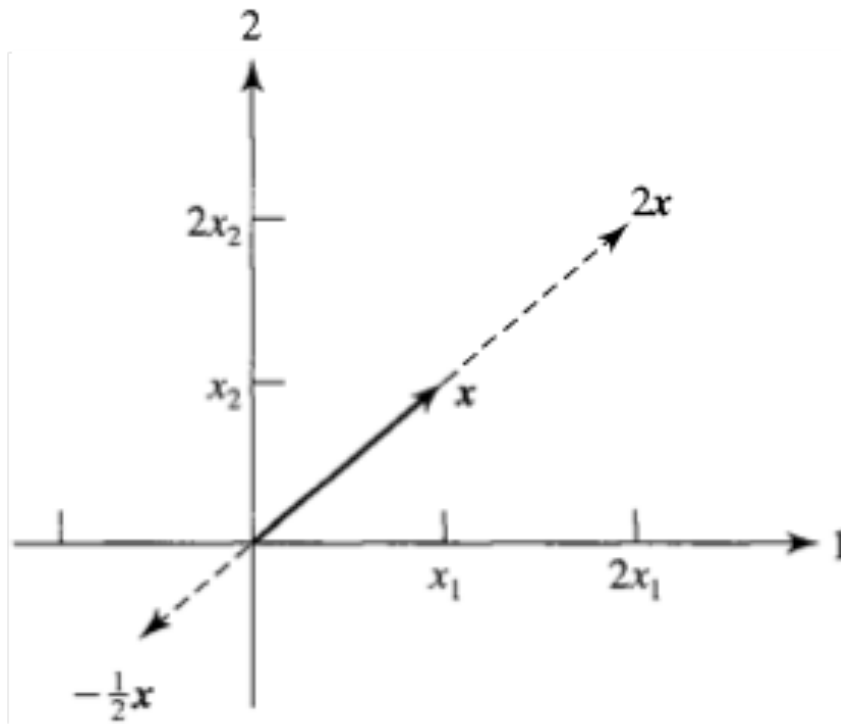
<https://stats.stackexchange.com/questions/2691/making-sense-of-principal-component-analysis-eigenvectors-eigenvalues>

# Points and Vectors

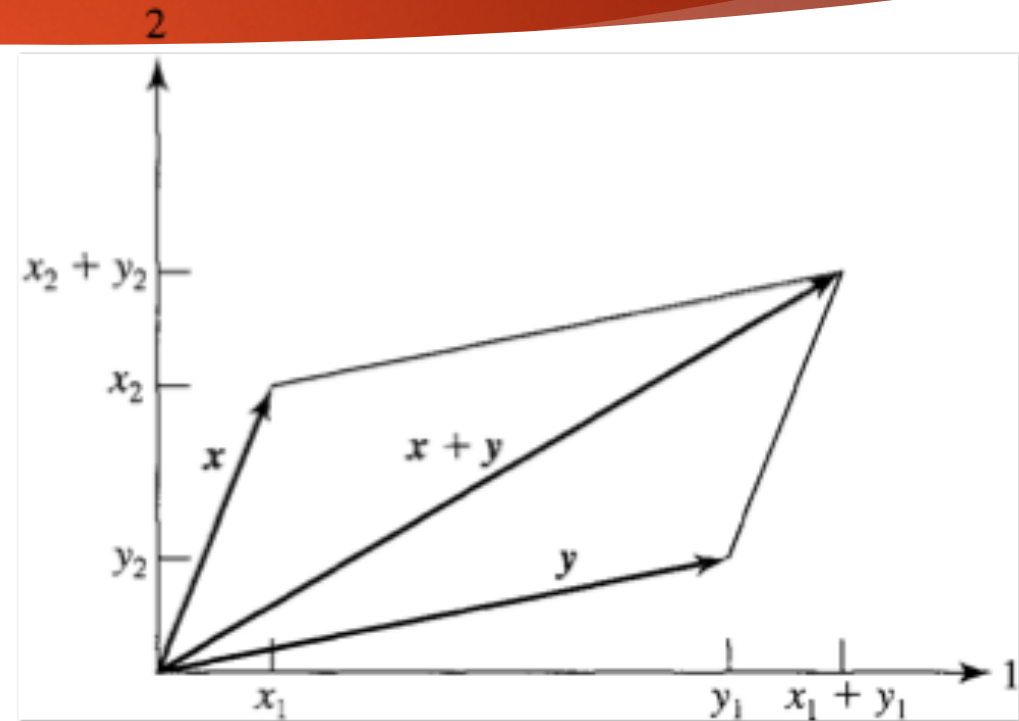


- ▶ Every point can be thought of a vector from the origin to that point
- ▶  $p = (1, 3, 2)$

# Scalar Multiplication and Vector Addition



(a)



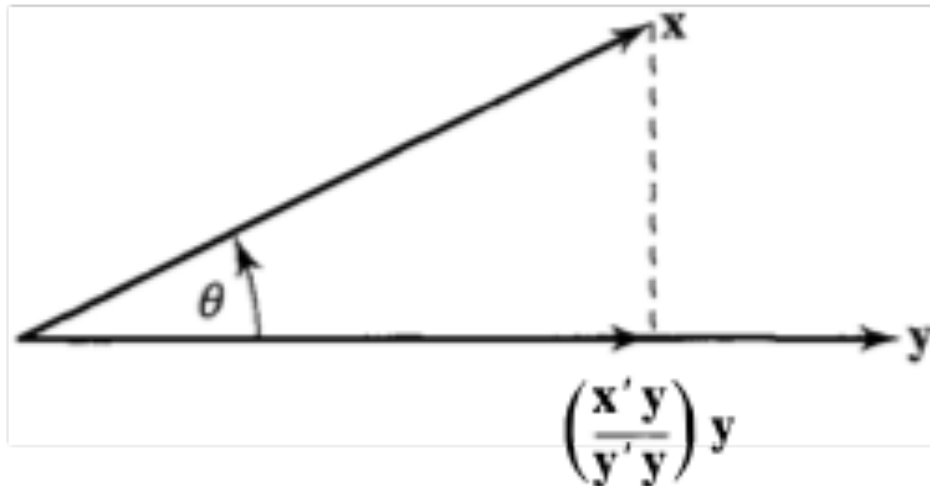
(b)



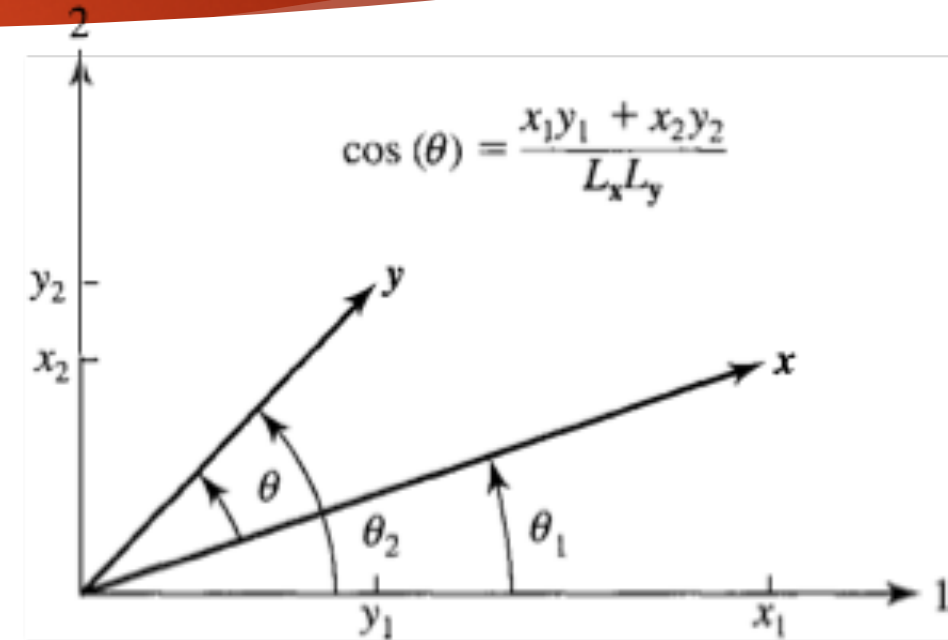
# Dot Product, Angles, Projections

$$\mathbf{x}'\mathbf{y} = x_1y_1 + x_2y_2 + \cdots + x_ny_n$$

$$\text{Projection of } \mathbf{x} \text{ on } \mathbf{y} = \frac{(\mathbf{x}'\mathbf{y})}{\mathbf{y}'\mathbf{y}} \mathbf{y} = \frac{(\mathbf{x}'\mathbf{y})}{L_y} \frac{1}{L_y} \mathbf{y}$$

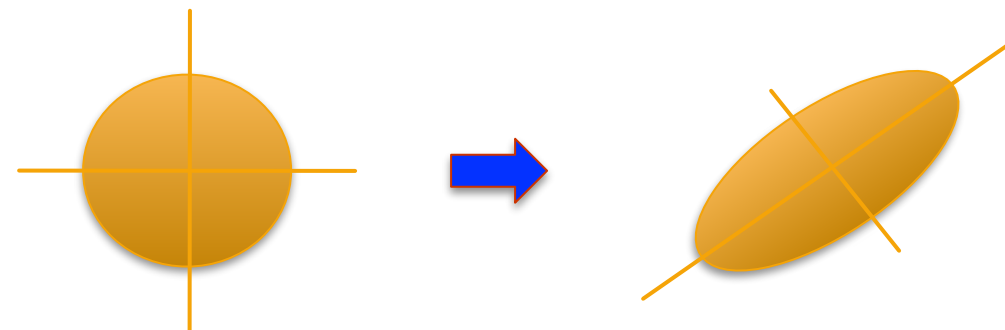


$$\leftarrow L_x \cos(\theta) \rightarrow$$



# Matrices & Transformations

- ▶ Arrays of Values,  $A$
- ▶ Linear Transformations
  - ▣  $Ax = y$
- ▶ Matrix Product
  - ▣ Composing transforms
- ▶ Matrix Inverse:  $AB = I \rightarrow B = A^{-1}$



# Data as Matrices

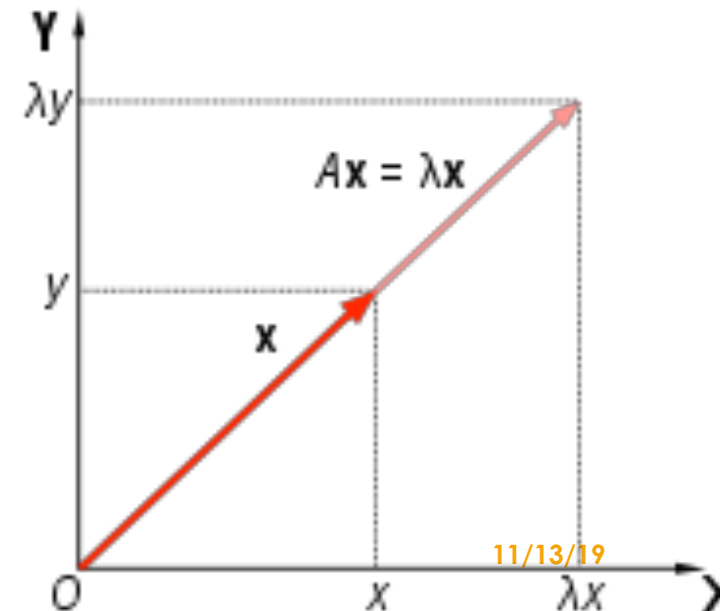
$$\mathbf{X}_{(n \times p)} = \begin{bmatrix} x_{11} & x_{12} & \cdots & x_{1p} \\ x_{21} & x_{22} & \cdots & x_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ x_{n1} & x_{n2} & \cdots & x_{np} \end{bmatrix} = \begin{bmatrix} \mathbf{x}'_1 \\ \mathbf{x}'_2 \\ \vdots \\ \mathbf{x}'_n \end{bmatrix}$$

← 1st (multivariate) observation

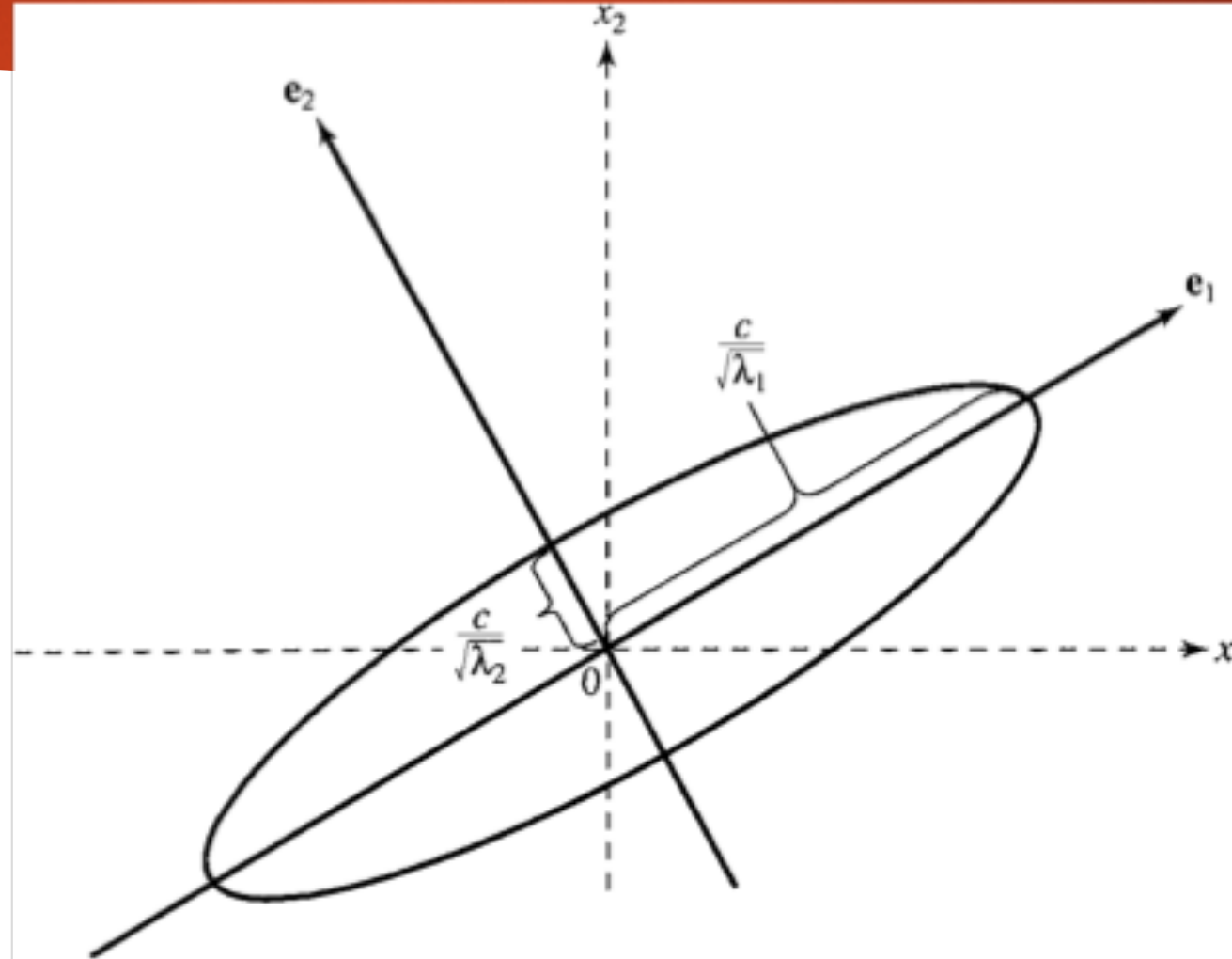
←  $n$ th (multivariate) observation

# Eigenvalues and Eigenvectors

- ▶ Under transform  $A$ , eigenvectors experience change in magnitude only, but not direction
- ▶  $Ax = \lambda x$ ;  $(A - \lambda I)x = 0$
- ▶ Characteristic Eq:  $|A - \lambda I| = 0$
- ▶ Eigenvalues:  $\lambda$
- ▶ Eigenvectors:  $x, e$



# Eigenvalues and Eigenvectors



# Spectral Decomposition

- ▶ If  $A$  is symmetric, then the following decomposition holds true:

$$\mathbf{A} = \lambda_1 \mathbf{e}_1 \mathbf{e}_1' + \lambda_2 \mathbf{e}_2 \mathbf{e}_2' + \dots + \lambda_k \mathbf{e}_k \mathbf{e}_k'$$

$(k \times k)$        $(k \times 1)(1 \times k)$        $(k \times 1)(1 \times k)$        $(k \times 1)(1 \times k)$

# Quadratic Form

- ▶ The scalar  $x'Ax$  is called **quadratic form**
- ▶ A is **positive definite**
  - ▣ if  $x'Ax > 0$ , whenever  $x$  is a nonzero vector
- ▶ Equivalently, A is **positive definite**
  - ▣ if all its eigenvalues are positive

# Matrix Inverse & Square Root

$$\mathbf{A} = \sum_{i=1}^k \lambda_i \mathbf{e}_i \mathbf{e}_i' = \mathbf{P} \mathbf{\Lambda} \mathbf{P}'$$

$(k \times k)$        $(k \times 1)(1 \times k)$        $(k \times k)(k \times k)(k \times k)$

$$\mathbf{A}^{-1} = \mathbf{P} \mathbf{\Lambda}^{-1} \mathbf{P}' = \sum_{i=1}^k \frac{1}{\lambda_i} \mathbf{e}_i \mathbf{e}_i'$$

$$\mathbf{A}^{1/2} = \sum_{i=1}^k \sqrt{\lambda_i} \mathbf{e}_i \mathbf{e}_i' = \mathbf{P} \mathbf{\Lambda}^{1/2} \mathbf{P}'$$

$$\mathbf{P} = [\mathbf{e}_1, \mathbf{e}_2, \dots, \mathbf{e}_k]$$

$$\mathbf{\Lambda} = \begin{bmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_k \end{bmatrix}$$

$(k \times k)$



# Dimension Reduction Revisited

► If we take  $r$  eigenvectors, then

▣  $P_r = [e_1, e_2, \dots, e_r]$ , and

▣  $A$  can be approximated by taking  $r$  eigenvectors

$$\Lambda_{(r \times r)} =$$

$$\begin{bmatrix} \lambda_1 & 0 & \dots & 0 \\ 0 & \lambda_2 & \dots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \dots & \lambda_r \end{bmatrix}$$

$$\begin{matrix} P & \Lambda & P' \\ (k \times r) & (r \times r) & (r \times k) \end{matrix}$$

# Random Matrices

$$E(\mathbf{X}) = \begin{bmatrix} E(X_{11}) & E(X_{12}) & \cdots & E(X_{1p}) \\ E(X_{21}) & E(X_{22}) & \cdots & E(X_{2p}) \\ \vdots & \vdots & \ddots & \vdots \\ E(X_{n1}) & E(X_{n2}) & \cdots & E(X_{np}) \end{bmatrix}$$

$$E(\mathbf{X} + \mathbf{Y}) = E(\mathbf{X}) + E(\mathbf{Y})$$
$$E(\mathbf{AXB}) = \mathbf{A}E(\mathbf{X})\mathbf{B}$$

# Covariance Matrix

$$\Sigma = E(\mathbf{X} - \boldsymbol{\mu})(\mathbf{X} - \boldsymbol{\mu})'$$

$$\Sigma = \text{Cov}(\mathbf{X}) = \begin{bmatrix} \sigma_{11} & \sigma_{12} & \cdots & \sigma_{1p} \\ \sigma_{21} & \sigma_{22} & \cdots & \sigma_{2p} \\ \vdots & \vdots & \ddots & \vdots \\ \sigma_{p1} & \sigma_{p2} & \cdots & \sigma_{pp} \end{bmatrix}$$

# Correlation Matrix, $\rho$

$$\mathbf{V}^{1/2} = \begin{bmatrix} \sqrt{\sigma_{11}} & 0 & \cdots & 0 \\ 0 & \sqrt{\sigma_{22}} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \sqrt{\sigma_{pp}} \end{bmatrix}$$

$$\boldsymbol{\rho} = (\mathbf{V}^{1/2})^{-1} \boldsymbol{\Sigma} (\mathbf{V}^{1/2})^{-1}$$

$$\mathbf{V}^{1/2} \boldsymbol{\rho} \mathbf{V}^{1/2} = \boldsymbol{\Sigma}$$

# Singular Value Decomposition

- ▶ Spectral Decomp. for sq. symm. matrices
- ▶ Non-sq. asymmetric matrices?
  - ▣ Use sq. root of eigenvalues of  $AA'$
  - ▣ Singular values of  $A$

$$P \quad \Lambda \quad P'$$

$$(k \times r) \quad (r \times r) \quad (r \times k)$$

$$A_{(m \times k)} = U_{(m \times m)} \Lambda_{(m \times k)} V'_{(k \times k)}$$

# Dimensionality Reduction

- ▶ Given  $m \times k$  matrix  $A$ , we can approximate it by  $m \times s$  matrix  $B$  with  $s < k = \text{rank}(A)$ . Then

$$\mathbf{B} = \sum_{i=1}^s \lambda_i \mathbf{u}_i \mathbf{v}_i'$$

- ▶ Here we are picking  $s$  singular values from SVD